

Exercise 8.5

Answer 1E.

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(a)

Here, $f(x)$ denote the probability density function for the life time of a manufactures highest quality car tire. The variable x is measured in miles.

The probability density function for the life time of a manufactures highest quality car tire $\mu = 40,000$ and $\sigma = 30,000$ miles is defined as,

$$\int_{30,000}^{40,000} f(x) dx .$$

Hence, the integral $\int_{30,000}^{40,000} f(x) dx$ is the probability density function for the life time of a manufactures highest quality car tire.

(b)

Here, the probability density function for the life time of a manufactures highest quality car tire which have life time more than or equal to 25,000 miles is denoted as,

$$\int_{25,000}^{\infty} f(x) dx .$$

Hence, the integral $\int_{25,000}^{\infty} f(x) dx$ is the probability density function for the life time of a manufactures highest quality car tire which have a life time more than or equal to 25,000 miles.

Answer 2E.

(A) The probability that we drive to school in less than 15 minutes is $\int_0^{15} f(t) dt$

(B) The probability that we drive to school in more than half an hour is $\int_{30}^{\infty} f(t) dt$

Answer 3E.

If $\int_{-\infty}^{\infty} f(x) dx = 1$, then $f(x)$ is a probability density function.

Keeping this in view, we are given that $f(x) = 30x^2(1-x)^2$ for $0 \leq x \leq 1$ and
 $f(x) = 0$ for all other values of x .

$$\begin{aligned}\text{Consider } \int_{-\infty}^{\infty} f(x) dx &= \int_{-\infty}^0 f(x) dx + \int_0^1 f(x) dx + \int_1^{\infty} f(x) dx \text{ keeping the interval in view.} \\ &= 0 + \int_0^1 30x^2(1-x)^2 dx + 0 \\ &= 30 \int_0^1 x^2(1+x^2-2x) dx \\ &= 30 \int_0^1 (x^2+x^4-2x^3) dx \\ &= 30 \left[\frac{x^3}{3} + \frac{x^5}{5} - 2 \frac{x^4}{4} \right]_0^1 \\ &= 30 \left[\frac{1}{3} + \frac{1}{5} - \frac{1}{2} \right] \\ &= 30 \left[\frac{10+6-15}{30} \right] \\ &= 1 \\ \int_{-\infty}^{\infty} f(x) dx &= 1\end{aligned}$$

$\therefore f$ is a probability density function

$$\begin{aligned}\text{(b)} \quad p\left(X \leq \frac{1}{3}\right) &= \int_{-\infty}^{1/3} f(x) dx \\ &= \int_{-\infty}^0 f(x) dx + \int_0^{1/3} f(x) dx \\ &= \int_0^{1/3} 30x^2(1-x)^2 dx \\ &= 30 \int_0^{1/3} (x^2+x^4-2x^3) dx \\ &= 30 \left[\frac{x^3}{3} + \frac{x^5}{5} - \frac{x^4}{2} \right]_0^{1/3} \\ &= 30 \left[\frac{1}{81} + \frac{1}{5} \cdot \frac{1}{35} - \frac{1}{2} \cdot \frac{1}{34} \right] \\ &= \frac{17}{81}\end{aligned}$$

$$\begin{aligned}\text{Therefore } p\left(X \leq \frac{1}{3}\right) &= \frac{17}{81} \\ &\approx 0.2098\end{aligned}$$

Answer 4E.

Consider the function,

$$f(x) = \begin{cases} xe^{-x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

(a)

The objective is to verify f is a probability density function or not.

To verify that f is probability density function, it is enough to show that $\int_{-\infty}^{\infty} f(x) dx = 1$

consider $\int_{-\infty}^{\infty} f(x) dx$

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \int_{-\infty}^0 f(x) dx + \int_0^{\infty} f(x) dx \\ &= 0 + \int_0^{\infty} xe^{-x} dx \\ &= \lim_{t \rightarrow \infty} \int_0^t xe^{-x} dx \end{aligned}$$

Now find the integration by using the following formula.

$$\int u dv = uv - \int v du$$

Here $u = x$, $dv = e^{-x} dx$

$$du = dx, v = -e^{-x}$$

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \lim_{t \rightarrow \infty} \left[x(-e^{-x}) - \int (-e^{-x}) dx \right] \\ &= \lim_{t \rightarrow \infty} \left[x(-e^{-x}) - (e^{-x}) \right]_0^t \\ &= \lim_{t \rightarrow \infty} \left[-te^{-t} - e^{-t} + 1 \right] \end{aligned}$$

Here, if $t \rightarrow \infty$ then $e^{-t} \rightarrow 0$

So

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \lim_{t \rightarrow \infty} (-te^{-t}) - \lim_{t \rightarrow \infty} (e^{-t}) + \lim_{t \rightarrow \infty} (1) \\ &= 0 - 0 + 1 \\ &= 1 \end{aligned}$$

Therefore, the value of the integral is,

$$\int_{-\infty}^{\infty} f(x) dx = \boxed{1}$$

Hence, the function $f(x)$ is the probability density function.

(b)

The objective is to find the value of $P(1 \leq X \leq 2)$

$$\begin{aligned}P(1 \leq X \leq 2) &= \int_1^2 f(x) dx \\&= \int_1^2 x e^{-x} dx \\&= \left[x(-e^{-x}) - (e^{-x}) \right]_1^2 \\&= \left[(-2e^{-2} - e^{-2}) - (-e^{-1} - e^{-1}) \right] \\&= \left(-\frac{2}{e^2} - \frac{1}{e^2} \right) - \left(-\frac{1}{e} - \frac{1}{e} \right) \\&= -\frac{3}{e^2} + \frac{2}{e} \\&= \frac{2}{e} - \frac{3}{e^2}\end{aligned}$$

Therefore, the value is,

$$P(1 \leq X \leq 2) = \boxed{\frac{2}{e} - \frac{3}{e^2}}.$$

Answer 5E.

$$\text{Given that } f(x) = \frac{c}{1+x^2}$$

(a)

We have to find that for what values of c is f a probability density function.

A function f is a probability density function if the following condition is satisfied.

$$\int_{-\infty}^{\infty} f(x) dx = 1$$

Therefore

$$\int_{-\infty}^{\infty} \frac{c}{1+x^2} dx = 1$$

$$\int_{-\infty}^{\infty} \frac{c}{1+x^2} dx = 1$$

$$c \left(\tan^{-1} x \right)_{-\infty}^{\infty} = 1$$

$$c \left(\frac{\pi}{2} + \frac{\pi}{2} \right) = 1$$

$$\boxed{c = \frac{1}{\pi}}$$

b)

Using the above value $c = \frac{1}{\pi}$, we have to find $P(-1 < X < 1)$.

This is done by evaluating the probability density function from -1 to 1:

$$\begin{aligned}&= \int_{-1}^1 f(x) dx \\&= \int_{-1}^1 \frac{c}{1+x^2} dx \\&= \int_{-1}^1 \frac{1/\pi}{1+x^2} dx \\&= \frac{1}{\pi} \left(\tan^{-1}(x) \right)_{-1}^1 \\&= \frac{1}{\pi} \left(\frac{\pi}{4} + \frac{\pi}{4} \right) \\&= \boxed{\frac{1}{2}}\end{aligned}$$

Answer 6E.

A function f is called a probability density function if it satisfies:

$$\int_{-\infty}^{\infty} f(x) dx = 1$$

The probability of a function f is defined as shown below:

$$P(a < X < b) = \int_a^b f(x) dx$$

The mean of a function is defined as shown below:

$$\mu = \int_{-\infty}^{\infty} xf(x) dx$$

Consider the function shown below:

$$f(x) = \begin{cases} 0, & x < 0 \\ k(3x - x^2), & 0 \leq x \leq 3 \\ 0, & x > 3 \end{cases}$$

a.

Evaluate the integral shown below:

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \int_{-\infty}^0 f(x) dx + \int_0^3 f(x) dx + \int_3^{\infty} f(x) dx \\ &= \int_{-\infty}^0 0 dx + \int_0^3 k(3x - x^2) dx + \int_3^{\infty} 0 dx \\ &= 0 + \int_0^3 k(3x - x^2) dx + 0 \\ &= k \int_0^3 3x dx - k \int_0^3 x^2 dx \end{aligned}$$

Evaluate the integral further:

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \frac{3k}{2} x^2 \Big|_0^3 - \frac{k}{3} x^3 \Big|_0^3 \\ &= \frac{3k}{2} (3^2 - 0^2) - \frac{k}{3} (3^3 - 0^3) \\ &= \frac{27k}{2} - \frac{27k}{3} \\ &= \frac{27k}{6} \end{aligned}$$

Determine the value of k :

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= 1 \\ \frac{27k}{6} &= 1 \\ k &= \frac{6}{27} \end{aligned}$$

Hence, the value of k is $\boxed{k = \frac{6}{27}}$.

b.

Substitute the value of k in the function:

$$f(x) = \begin{cases} 0, & x < 0 \\ \frac{6}{27}(3x - x^2), & 0 \leq x \leq 3 \\ 0, & x > 3 \end{cases}$$

Evaluate the probability of the function as shown below:

$$\begin{aligned} P(X > 1) &= \int_1^{\infty} f(x) dx \\ &= \frac{6}{27} \int_1^3 (3x - x^2) dx + \frac{6}{27} \int_3^{\infty} 0 dx \\ &= \frac{6}{27} \left[3 \frac{x^2}{2} - \frac{x^3}{3} \right]_1^3 + 0 \\ &= \frac{6}{27} \left[\left(\frac{27}{2} - \frac{27}{3} \right) - \left(\frac{3}{2} - \frac{1}{3} \right) \right] \end{aligned}$$

Continue further to obtain the value:

$$\begin{aligned} P(X > 1) &= \frac{6}{27} \left[\frac{27}{2} - \frac{27}{3} - \frac{3}{2} + \frac{1}{3} \right] \\ &= \frac{6}{27} \left[\frac{24}{2} - 9 + \frac{1}{3} \right] \\ &= \frac{20}{27} \end{aligned}$$

Hence, the value of the probability is $P(X > 1) = \frac{20}{27}$.

c.

Use the formula of mean to obtain the required value:

$$\begin{aligned} \mu &= \int_{-\infty}^{\infty} xf(x) dx \\ &= \int_{-\infty}^0 xf(x) dx + \int_0^3 x \frac{6}{27} (3x - x^2) dx + \int_3^{\infty} xf(x) dx \\ &= 0 + \frac{6}{27} \int_0^3 (3x^2 - x^3) dx + 0 \\ &= \frac{6}{27} \left(3 \times \frac{x^3}{3} - \frac{x^4}{4} \right)_0^3 \end{aligned}$$

Continue further to evaluate the integral:

$$\begin{aligned} \mu &= \frac{6}{27} \left(x^3 - \frac{x^4}{4} \right)_0^3 \\ &= \frac{6}{27} \left[\left(3^3 - \frac{3^4}{4} \right) - \left(0^3 - \frac{0^4}{4} \right) \right] \\ &= \frac{6}{27} \left(27 - \frac{81}{4} \right) \\ &= \frac{3}{2} \end{aligned}$$

So, the mean is 1.5.

Hence, the final value of the mean is $\mu = 1.5$.

Answer 7E.

(A) We have the function $f(x) = \begin{cases} 0.1 & \text{if } 0 \leq x \leq 10 \\ 0 & \text{if } x < 0 \text{ or } x > 10 \end{cases}$

We see that in the interval $0 \leq x \leq 10$, $f(x) \geq 0$

And for $x < 0$ or $x > 10$, $f(x) = 0$

So $f(x) \geq 0$ for all x

Now we have to check $\int_{-\infty}^{\infty} f(x) dx = 1$

$$\begin{aligned} \Rightarrow \int_{-\infty}^{\infty} f(x) dx &= \int_0^{10} 0.1 dx \\ &= 0.1 [x]_0^{10} \\ &= 1 \end{aligned}$$

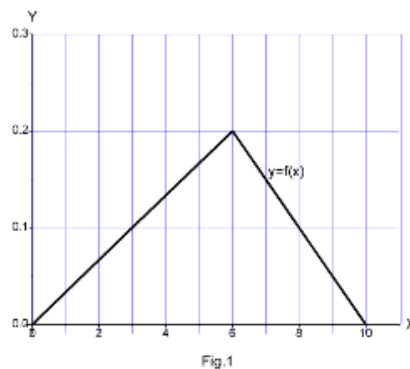
So $f(x)$ is a probability density function

(B) The value of mean will be the mid point of the interval, which is 5

Now we calculate the mean

$$\begin{aligned} \mu &= \int_{-\infty}^{\infty} x f(x) dx = \int_0^{10} 0.1x dx \\ &= 0.1 \left[\frac{x^2}{2} \right]_0^{10} \\ \Rightarrow \mu &= 0.1 [50] \\ &= 5 \\ \boxed{\mu = 5} \end{aligned}$$

Answer 8E.



(A) Here we see from the graph that $f(x) \geq 0$ for all x

Now we check $\int_{-\infty}^{\infty} f(x) dx = 1$

$$\int_{-\infty}^{\infty} f(x) dx = \int_0^{10} f(x) dx = \text{Area of the region from } x=0 \text{ to } x=10$$

This is the area of a triangle

$$\begin{aligned} \text{So } \int_{-\infty}^{\infty} f(x) dx &= \frac{1}{2} \times 10 \times 0.2 \\ &= 1 \end{aligned}$$

So $f(x)$ is a probability density function

(B) (1) $P(x < 3) = \int_0^3 f(x) dx$

$$\begin{aligned} &= \text{Area of the triangle from } x=0 \text{ to } x=3 \\ &= \frac{1}{2} \times 3 \times 0.1 \\ \boxed{P(x < 3) = 0.15} \end{aligned}$$

(2) We have to find $P(3 \leq x \leq 8)$

$$\begin{aligned}
 P(3 \leq x \leq 8) &= \int_3^8 f(x) dx \\
 &= \int_0^8 f(x) dx - \int_0^3 f(x) dx \\
 &= \int_0^{10} f(x) dx - \int_8^{10} f(x) dx - \int_0^3 f(x) dx \\
 &= 1 - \frac{1}{2} \times 0.1 \times 2 - 0.15 \\
 &= 1 - 0.25 \\
 &= 0.75
 \end{aligned}$$

$$\boxed{P(3 \leq x \leq 8) = 0.75}$$

(C) Mean $\int_{-\infty}^{\infty} x f(x) dx = \int_0^{10} x f(x) dx$

We divide the interval $[0, 10]$ in to 10 subintervals

$$\begin{aligned}
 \text{Then } \Delta x &= \frac{10-0}{10} \\
 &= 1 \\
 \Rightarrow \Delta x &= 1
 \end{aligned}$$

Then by Simpson's rule we can approximate the value of mean

$$\mu = \int x f(x) dx \approx \frac{\Delta x}{3} [x_0 f(x_0) + 4x_1 f(x_1) + 2x_2 f(x_2) + 4x_3 f(x_3) + \dots + 4x_9 f(x_9) + x_{10} f(x_{10})]$$

$$\approx \frac{1}{3} [0 + 4.(1).(0) + 2.(2).(0.04) + 4.(3).(0.08) + 2.(4).(0.12) + 4.(5).(0.16) + 2.(6).(0.2) + 4.(7).(0.15) + 2.(8).(0.1) + 4.(9).(0.05) + 10.(0)]$$

$$\approx \frac{1}{3} [0.16 + 0.96 + 0.96 + 3.2 + 2.4 + 4.2 + 1.6 + 1.8]$$

Or $\boxed{\mu \approx 5.09}$

Or $\boxed{\mu = 5}$

Answer 9E.

We have the probability density function

$$f(t) = \begin{cases} 0 & \text{if } t < 0 \\ 0.2e^{-t/5} & \text{if } t \geq 0 \end{cases}$$

If m is the median waiting time for a phone call to the company then

$$\begin{aligned}
 \int_m^{\infty} f(t) dt &= \frac{1}{2} \\
 \Rightarrow \int_m^{\infty} 0.2e^{-t/5} dt &= \frac{1}{2} \\
 \Rightarrow \lim_{x \rightarrow \infty} \int_m^x 0.2e^{-t/5} dt &= \frac{1}{2} \\
 \Rightarrow \lim_{x \rightarrow \infty} [-5 \times 0.2 \times e^{-t/5}]_m^x &= \frac{1}{2}
 \end{aligned}$$

Therefore

$$\Rightarrow \lim_{x \rightarrow \infty} [-e^{-x/5} + e^{-m/5}] = \frac{1}{2}$$

$$\Rightarrow e^{-m/5} - \lim_{x \rightarrow \infty} e^{-x/5} = \frac{1}{2}$$

$$\Rightarrow e^{-m/5} = 1/2$$

$$\left[\lim_{x \rightarrow \infty} e^{-x/5} = 0 \right]$$

$$\Rightarrow \frac{-m}{5} = \ln(1/2)$$

$$\Rightarrow \frac{-m}{5} = -\ln 2 \Rightarrow \boxed{m = 5 \ln 2} \text{ min}$$

$$\Rightarrow \boxed{m \approx 3.47} \text{ min}$$

Answer 10E.

(A) We have exponential density function

$$f(t) = \begin{cases} 0 & \text{if } t < 0 \\ \mu^{-1}e^{-t/\mu} & \text{if } t \geq 0 \end{cases}$$

We have $\mu = 1000$

$$\text{Then } f(t) = \begin{cases} 0 & \text{if } t < 0 \\ (0.001)e^{-0.001t} & \text{if } t \geq 0 \end{cases}$$

(1) We have to find $P(t \leq 200)$

$$\begin{aligned} \Rightarrow P(t < 200) &= \int_0^{200} (0.001)e^{-0.001t} dt \\ &= \left[\frac{0.001}{-0.001} e^{-0.001t} \right]_0^{200} \\ &= \left[-e^{(0.001 \times 200)} + e^0 \right] \\ &= 1 - e^{-0.2} \approx 0.18 \end{aligned}$$

$$\boxed{P(t \leq 200) \approx 0.18}$$

(2) We have to find $P(t > 800)$

$$\begin{aligned} P(t > 800) &= \int_{800}^{\infty} (0.001)e^{-0.001t} dt \\ &= \lim_{x \rightarrow \infty} \left[-e^{-0.001t} \right]_{800}^x \\ &= \lim_{x \rightarrow \infty} \left[-e^{-(0.001)x} + e^{-0.8} \right] \\ &= e^{-0.8} \end{aligned}$$

$$\left[\lim_{x \rightarrow \infty} e^{-(0.001)x} = 0 \right]$$

$$\boxed{P(t > 800) \approx 0.45}$$

(B) If median is m , then $\int_m^{\infty} f(t) dt = \frac{1}{2}$

$$\begin{aligned} \Rightarrow \int_m^{\infty} (0.001)e^{-0.001t} dt &= \frac{1}{2} \\ \Rightarrow \lim_{x \rightarrow \infty} \int_m^x (0.001)e^{-0.001t} dt &= \frac{1}{2} \\ \Rightarrow \lim_{x \rightarrow \infty} \left[-e^{-0.001t} \right]_m^x &= \frac{1}{2} \end{aligned}$$

Therefore

$$\begin{aligned} \Rightarrow \lim_{x \rightarrow \infty} \left[e^{-0.001m} - e^{-0.001x} \right] &= \frac{1}{2} \\ \Rightarrow e^{-0.001m} &= \frac{1}{2} \\ \Rightarrow -0.001m &= \ln\left(\frac{1}{2}\right) \\ \Rightarrow -0.001m &= -\ln 2 \\ \Rightarrow 0.001m &= \ln 2 \\ \Rightarrow m &= \frac{1}{0.001} \ln 2 \\ \Rightarrow \boxed{m = 1000 \ln 2} &\approx 693.15 \text{ hours} \end{aligned}$$

Answer 11E.

(A) We have $\mu = 2.5$ minutes

Thus probability density function as

$$f(t) = \begin{cases} 0 & \text{if } t < 0 \\ \frac{1}{2.5} e^{-t/2.5} & \text{if } t \geq 0 \end{cases}$$

Then probability that a customer has to wait for more than 4 minutes is

$$\begin{aligned}
 P(t > 4) &= \int_4^{\infty} \frac{1}{2.5} e^{-t/2.5} dt \\
 &= \lim_{x \rightarrow \infty} \int_4^x \frac{1}{2.5} e^{-t/2.5} dt \\
 &= \lim_{x \rightarrow \infty} \left[-e^{-t/2.5} \right]_4^x \\
 &= \lim_{x \rightarrow \infty} \left[-e^{-x/2.5} + e^{-4/2.5} \right] \\
 &= e^{-4/2.5} \approx 0.20
 \end{aligned}$$

- (B) Probability that a customer is served with in first 2 minutes is

$$\begin{aligned}
 P(t \leq 2) &= \int_0^2 \frac{1}{2.5} e^{-t/2.5} dt \\
 &= \left[-e^{-t/2.5} \right]_0^2 \\
 &= \left[-e^{-2/2.5} + e^0 \right] = 1 - e^{-0.8} \\
 &\approx 0.55 \\
 \boxed{P(t \leq 2) \approx 0.55}
 \end{aligned}$$

- (C) Let the customer has been given free hamburgers after 'a' minutes

$$\begin{aligned}
 \text{Then } \int_a^{\infty} \frac{1}{2.5} e^{-t/2.5} dt &= 0.02 \\
 \Rightarrow \lim_{x \rightarrow \infty} \left[-e^{-t/2.5} \right]_a^x &= 0.02 \\
 \Rightarrow \lim_{x \rightarrow \infty} \left[e^{-x/2.5} + e^{-a/2.5} \right] &= 0.02 \\
 \Rightarrow e^{-a/2.5} &= 0.02 \\
 \Rightarrow \frac{-a}{2.5} &= \ln(0.02) \\
 \Rightarrow a &= -2.5 \times \ln(0.02) \\
 \Rightarrow a &\approx 9.78 \text{ minutes} = \boxed{10 \text{ minutes}}
 \end{aligned}$$

Thus the advertisement should say that if the customer is not served with in 10 minutes, he get a free hamburger.

Answer 12E.

- (A) We have $\mu = 69$ and $\sigma = 2.8$

$$\text{Then probability density function is } f(x) = \frac{1}{2.8\sqrt{2\pi}} e^{-(x-69)^2 / 2(2.8)^2}$$

And the probability that an adult male chosen at random is between 65 inches and 73 inches tall is

$$P(65 \leq x \leq 73) = \int_{65}^{73} \frac{1}{2.8\sqrt{2\pi}} e^{-(x-69)^2 / 2(2.8)^2} dx$$

Now we use Simpson's rule to evaluate this integral interval is $[65, 73]$

$$\text{Taking } n = 8 \Rightarrow \Delta x = (73 - 65) / 8 = 1$$

Subintervals are $[65, 66], [66, 67], [67, 68], [68, 69], [69, 70], [70, 71], [71, 72], [72, 73]$

Then by Simpson's rule

$$\begin{aligned}
 P(65 \leq x \leq 73) &\approx \frac{1}{2.8\sqrt{2\pi}} \cdot \frac{1}{3} \left[e^{-(65-69)^2 / 2(2.8)^2} + 4e^{-(66-69)^2 / 2(2.8)^2} + 2e^{-(67-69)^2 / 2(2.8)^2} \right. \\
 &\quad \left. \dots + 4e^{-(72-69)^2 / 2(2.8)^2} + e^{-(73-69)^2 / 2(2.8)^2} \right] \\
 &\approx 0.8469 \\
 \Rightarrow \boxed{P(65 \leq x \leq 73) = 85\%}
 \end{aligned}$$

- (B) The probability that the adult male population is more than 6feet = $6 \times 12 = 72$ inches tall (1 feet = 12 inch)

$$P(x > 72) = 1 - \int_0^{72} \frac{1}{2.8\sqrt{2\pi}} e^{-(x-69)^2 / 2(2.8)^2} dx$$

For evaluating this integral we use Simpson's rule with $n = 36$

$$\text{Then } \Delta x = \frac{72-0}{36} = 2$$

Then subintervals are $[0, 2], [2, 4], \dots, [70, 72]$

$$\text{Let } g(x) = e^{-(x-69)^2/2(28)^2}$$

Then by Simpson's rule

$$P(0 \leq x \leq 72) = \frac{1}{2.8\sqrt{2\pi}} \frac{2}{3} [g(0) + 4g(2) + 2g(4) + 4g(6) + \dots + 4g(70) + g(72)]$$

$$\approx 0.858$$

$$\text{Then } P(x > 72) \approx 1 - 0.858$$

$$\Rightarrow \boxed{P(x > 72) \approx 14.2\%}$$

Answer 13E.

We have $\mu = 9.4$ and $\sigma = 4.2$

$$\text{Then probability density function is } f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-(x-\mu)^2/2(\sigma)^2}$$

$$f(x) = \frac{1}{4.2\sqrt{2\pi}} e^{-(x-9.4)^2/2(4.2)^2}$$

The probability of households throw out at least 10lb of paper a week is

$$P(x \geq 10) = \int_{10}^{\infty} f(x) dx$$

This is an improper integral so we can approximate it by the integral 10 to 100 because more than 100lb of paper extremely rare.

$$\text{Then } P(x > 10) = \int_{10}^{100} f(x) dx$$

$$\text{Where } f(x) = \frac{1}{4.2\sqrt{2\pi}} e^{-(x-9.4)^2/2(4.2)^2}$$

MIDPOINT RULE:

$$\int_a^b f(x) dx \approx M_n \approx \Delta x [f(\bar{x}_1) + 2f(\bar{x}_2) + 2f(\bar{x}_3) + \dots + f(\bar{x}_n)]$$

$$\text{where } \Delta x = \frac{b-a}{n} \text{ and } \bar{x}_i = \frac{1}{2}(x_{i-1} + x_i) = \text{midpoint of } [x_{i-1}, x_i]$$

We can use mid point rule for approximating this integral taking $n = 45$

$$\Delta x = \frac{100-10}{45} = 2$$

Then subintervals are $[10, 12], [12, 14], [14, 16], \dots, [98, 100]$

And mid points are 11, 13, 15, 17, 19, 21, 23, ..., 99

$$\text{Then } P(x \geq 10) \approx 10 \times [f(11) + f(13) + f(15) + f(17) + f(19) + f(21) + \dots + f(97) + f(99)]$$

$$\approx 0.44$$

$$\Rightarrow \boxed{P(x \geq 10) \approx 44\%}$$

Answer 15E.

Consider the speeds of vehicles on a highway with speed limit 100 km/h are normally distributed with mean 112 km/h and standard deviation 8 km/h.

(a)

Find the probability that a randomly chosen vehicle is travelling at a legal speed.

The probability density function is, $f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$

The legal speed is, 0 to 100 km/h, mean $\mu = 112$, standard deviation $\sigma = 8$.

To find the probability that a randomly chosen vehicle is travelling at a legal speed, find

$$\int_0^{100} f(x) dx.$$

Substitute mean $\mu = 112$, standard deviation $\sigma = 8$.

Therefore,

$$\begin{aligned}\int_0^{100} f(x) dx &= \int_0^{100} \frac{1}{8\sqrt{2\pi}} e^{-\frac{(x-112)^2}{2(8)^2}} dx \\ &= \int_0^{100} \frac{1}{8\sqrt{2\pi}} e^{-\frac{(x-112)^2}{128}} dx \\ &\approx 0.06681\end{aligned}$$

Thus, the probability that a randomly chosen vehicle is travelling at a legal speed is,

$$\boxed{0.06681}.$$

(b)

If police are instructed to ticket motorists driving 125 km/h or more, find the percentage of motorists are targeted.

To find the percentage of motorists is targeted, find the probability of $\int_{125}^{\infty} f(x) dx$.

Therefore,

$$\begin{aligned}\int_{125}^{\infty} f(x) dx &= \int_{125}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2(\sigma)^2}} dx \\ &= \int_{125}^{\infty} \frac{1}{8\sqrt{2\pi}} e^{-\frac{(x-112)^2}{2(8)^2}} dx \quad \text{Substitute } \sigma = 8, \mu = 112 \\ &= \int_{125}^{\infty} \frac{1}{8\sqrt{2\pi}} e^{-\frac{(x-112)^2}{128}} dx \\ &\approx 0.0521\end{aligned}$$

The percentage is, $(0.0521) \times 100 = \boxed{5.21\%}$.

Answer 16E.

The objective is to show that the probability distribution function for a normally distributed random variable has inflection points at $x = \mu \pm \sigma$.

Consider the normal distribution function,

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

Where $\mu > 0$ and $\sigma > 0$ called the standard deviation.

Finding the derivative $f'(x)$ we obtain

$$f'(x) = -\frac{(x-\mu)}{\sigma^3\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

Thus, $x = \mu$ is the only critical number. Finding the second derivative we obtain

$$f''(x) = \frac{1}{\sigma^3\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \left(\frac{(x-\mu)^2}{\sigma^2} - 1 \right)$$

Hence, $f''(\mu) = -\frac{1}{\sigma^3\sqrt{2\pi}} < 0$ so that $x = \mu$ is a maximum.

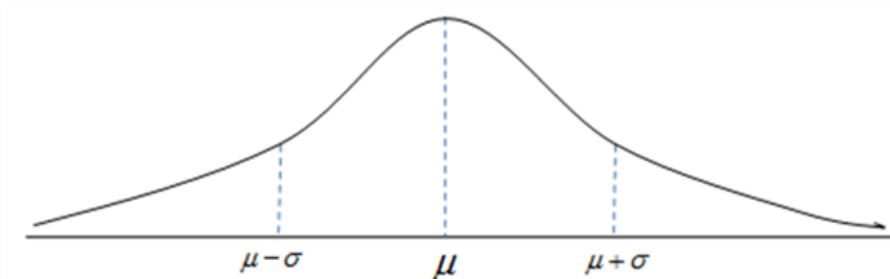
Solving the equation $f''(x) = 0$ we find $x = \mu \pm \sigma$.

If $\mu - \sigma < x < \mu + \sigma$ then $|x - \mu| < \sigma$ and therefore $\frac{(x-\mu)^2}{\sigma^2} < 1$ so that $f''(x) < 0$ and the graph of f is concave down.

Similarly, we see that $f''(x) > 0$ for $x < \mu - \sigma$ or $x > \mu + \sigma$ so that the graph of $f(x)$ is concave up.

It follows that $f(x)$ has points of inflection at $x = \mu \pm \sigma$.

The graph is shown below:



From the above figure, it is shown that the probability distribution function for a normally distributed random variable has inflection points at $x = \mu \pm \sigma$.

Answer 17E.

Two standard deviations of the mean are $\mu - 2\sigma$ and $\mu + 2\sigma$

$$\text{Then } P = \int_{\mu-2\sigma}^{\mu+2\sigma} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$\text{Let } \frac{x-\mu}{\sigma} = t$$

$$\Rightarrow \frac{1}{\sigma} dx = dt$$

And when $x = \mu - 2\sigma$, $\Rightarrow t = -2$ and when $x = \mu + 2\sigma$, $t = 2$

$$\text{Then } P = \int_{-2}^2 \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$$

$$\text{Let } f(t) = \frac{1}{\sqrt{2\pi}} e^{-t^2/2}$$

Interval is $[-2, 2]$

$$\text{Taking } n = 8, \quad \Delta x = \frac{2+2}{8} = \frac{1}{2}$$

Then by Simpson's rule

$$\begin{aligned}
 P &= \frac{\Delta x}{3} \left[f(-2) + 4f(-1.5) + 2f(-1) + 4f(-0.5) + 2f(0) + 4f(0.5) + 2f(1) \right. \\
 &\quad \left. + 4f(1.5) + f(2) \right] \\
 &= \frac{1}{6} \left[f(-2) + 4f(-1.5) + 2f(-1) + 4f(-0.5) + 2f(0) + 4f(0.5) \right. \\
 &\quad \left. + 2f(1) + 4f(1.5) + f(2) \right] \\
 &\Rightarrow \boxed{P \approx 0.9545}
 \end{aligned}$$

This is the probability that the random variable lies within two standard deviation of the mean.

Answer 18E.

Consider the standard deviation for a Random variable with probability density function f and μ is defined by

$$\sigma = \left[\int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx \right]^{\frac{1}{2}}$$

Need to find a standard deviation for an exponential density function with mean μ .

Write the probability density function as,

$$f(x) = \begin{cases} 0 & \text{if } t < 0 \\ \frac{1}{\mu} e^{-\frac{x}{\mu}} & \text{if } t \geq 0 \end{cases}$$

$$\text{That is } f(x) = \begin{cases} 0 & \text{if } -\infty < x < 0 \\ \frac{1}{\mu} e^{-\frac{x}{\mu}} & \text{if } 0 \leq x < \infty \end{cases}$$

Thus, the standard deviation

$$\sigma = \left[\int_{-\infty}^0 (x - \mu)^2 0 dx \right]^{\frac{1}{2}} + \left[\int_0^{\infty} (x - \mu)^2 \frac{1}{\mu} e^{-\frac{x}{\mu}} dx \right]^{\frac{1}{2}}$$

$$\sigma = 0 + \frac{1}{\sqrt{\mu}} \left[\int_0^{\infty} (x - \mu)^2 e^{-\frac{x}{\mu}} dx \right]^{\frac{1}{2}}$$

$$\sigma = \frac{1}{\sqrt{\mu}} \left[\int_0^{\infty} (x - \mu)^2 e^{-\frac{x}{\mu}} dx \right]^{\frac{1}{2}}$$

Squaring on both sides,

$$\begin{aligned}
 \sigma^2 &= \mu^2 \int_0^{\infty} \left(\frac{x}{\mu} - 1 \right)^2 e^{-\frac{x}{\mu}} dx \\
 &= \mu \lim_{x \rightarrow \infty} \int_0^x \left(\frac{x}{\mu} - 1 \right)^2 e^{-\frac{x}{\mu}} \frac{1}{\mu} dx
 \end{aligned}$$

Setting $\frac{x}{\mu} = t \Rightarrow \frac{1}{\mu} dx = dt$

$$\begin{aligned}
 \sigma^2 &= \mu \lim_{x \rightarrow \infty} \int_0^x (t-1)^2 e^{-t} dt \\
 &= \mu \lim_{x \rightarrow \infty} \left[(t-1)^2 \int e^{-t} dt - \int \frac{d(t-1)^2}{dt} \left(\int e^{-t} dt \right) dt \right] \\
 &= \mu \lim_{x \rightarrow \infty} \left[-(t-1)^2 e^{-t} \right]_0^x + \int 2(t-1) e^{-t} dt \\
 &= \mu \lim_{x \rightarrow \infty} \left[-(x-1)^2 e^{-x} + (0-1)^2 e^{-0} + 2(t-1) \int e^{-t} dt - \int \frac{d(t-1)}{dt} \left(\int e^{-t} dt \right) dt \right] \\
 &= \mu \lim_{x \rightarrow \infty} \left[-(x-1)^2 e^{-x} + 1 - [2(t-1)e^{-t}]_0^x + \int e^{-t} dt \right] \\
 &= \mu \lim_{x \rightarrow \infty} \left[-(x-1)^2 e^{-x} + 1 - 2(x-1)e^{-x} + 2(0-1)e^{-0} - [e^{-t}]_0^x \right] \\
 &= \mu \lim_{x \rightarrow \infty} \left[-(x-1)^2 e^{-x} + 1 - 2(x-1)e^{-x} + 2 - [e^{-x} - e^{-0}] \right] \\
 &= \mu \lim_{x \rightarrow \infty} \left[-(x-1)^2 e^{-x} + 1 - 2(x-1)e^{-x} + 2 - e^{-x} + 1 \right] \\
 &= \mu \left[-(\infty-1)^2 e^{-\infty} - 2(\infty-1)e^{-\infty} - e^{-\infty} + 4 \right] \\
 &= \mu \left[-(\infty-1)^2 (0) - 2(\infty-1)(0) - 0 + 4 \right] \\
 &= 4\mu \\
 \sigma^2 &= 4\mu
 \end{aligned}$$

Hence, the standard deviation for an exponential density function is $\boxed{\sigma = 16\mu^2}$.

Answer 19E.

The probability density function is

$$p(r) = \frac{4}{a_0^3} r^2 e^{-\frac{2r}{a_0}}, \quad r \geq 0$$

where a_0 is the Bohr radius ($a_0 \approx 5.59 \times 10^{-11}$ m).

Consider the integral

$$P(r) = \int_0^r \frac{4}{a_0^3} s^2 e^{-\frac{2s}{a_0}} ds$$

This integral gives the probability that the electron will be found within the sphere of radius r meters centered at the nucleus.

(a)

Need to verify that $p(r)$ is a probability density function.

$$\begin{aligned}\int_{-\infty}^{\infty} p(r) dr &= \int_0^{\infty} \left[\frac{4}{a_0^3} r^2 e^{-\frac{2r}{a_0}} \right] dr && [\text{For } r \geq 0] \\&= \frac{4}{a_0^3} \int_0^{\infty} \left(r^2 e^{-\frac{2r}{a_0}} \right) dr \\&= \frac{4}{a_0^3} \left[-\frac{1}{4} a_0 (a_0^2 + 2a_0 r + 2r^2) e^{-\frac{2r}{a_0}} \right]_0^{\infty} \\&= \frac{4}{a_0^3} \left[-\frac{1}{4} a_0 (a_0^2 + 2a_0 r + 2r^2) e^{-\infty} + \frac{1}{4} a_0 (a_0^2 + 2a_0 r + 2r^2) e^0 \right] \\&= \frac{4}{a_0^3} \left[-\frac{1}{4} a_0 (a_0^2 + 2a_0 r + 2r^2) (0) + \frac{1}{4} a_0 (a_0^2 + 2a_0 (0) + 2(0)^2) (1) \right] \\&= \frac{4}{a_0^3} \left[\frac{1}{4} a_0 (a_0^2) \right] \\&= 1\end{aligned}$$

Therefore, $\int_{-\infty}^{\infty} p(r) dr = 1$, so $p(r)$ is a probability density function.

(b) Find the limit of $p(r)$ as r tends to infinity.

$$\begin{aligned}\lim_{r \rightarrow \infty} p(r) &= \lim_{r \rightarrow \infty} \left(\frac{4}{a_0^3} r^2 e^{-\frac{2r}{a_0}} \right) \\&= 0 && \left[\text{Since as } r \rightarrow \infty, e^{-r} \rightarrow 0, e^{-\frac{2r}{a_0}} \rightarrow 0, r^2 e^{-\frac{2r}{a_0}} \rightarrow 0 \right]\end{aligned}$$

Consider $p(r) = \frac{4}{a_0^3} r^2 e^{-\frac{2r}{a_0}}$.

Differentiate with respect to r , get

$$\begin{aligned}p'(r) &= \frac{d}{dr} \left(\frac{4}{a_0^3} r^2 e^{-\frac{2r}{a_0}} \right) \\&= \frac{4}{a_0^3} \left(2r e^{-\frac{2r}{a_0}} - \frac{2}{a_0} r^2 e^{-\frac{2r}{a_0}} \right)\end{aligned}$$

Equate $p'(r)$ to zero and solve for the variable r , get

$$\frac{4}{a_0^3} \left(2re^{\frac{2r}{a_0}} - \frac{2}{k} r^2 e^{\frac{2r}{a_0}} \right) = 0$$

$$2re^{\frac{2r}{a_0}} - \frac{2}{k} r^2 e^{\frac{2r}{a_0}} = 0$$

$$re^{\frac{2r}{a_0}} \left(2 - \frac{2}{a_0} r \right) = 0$$

$$r \left(2 - \frac{2}{a_0} r \right) = 0$$

$$\text{either } r = 0 \text{ or } 2 - \frac{2}{a_0} r = 0$$

$$\text{either } r = 0 \text{ or } r = a_0$$

When $r = 0$,

$$p(r) = \frac{4}{a_0^3} r^2 e^{\frac{2r}{a_0}}$$

$$\begin{aligned} p(0) &= \frac{4}{a_0^3} (0)^2 e^{\frac{2(0)}{a_0}} \\ &= 0 \end{aligned}$$

When $r = a_0$,

$$p(r) = \frac{4}{a_0^3} r^2 e^{\frac{2r}{a_0}}$$

$$\begin{aligned} p(a_0) &= \frac{4}{a_0^3} (a_0)^2 e^{\frac{2(a_0)}{a_0}} \\ &= \frac{4}{a_0} e^{-2} \end{aligned}$$

Therefore $p(r)$ have its maximum value when $r = a_0 \approx 5.59 \times 10^{-11} \text{ m}$.

(c)

$$\text{Let } p(r) = \frac{4}{(5.59 \times 10^{-11})^3} r^2 e^{-\frac{2r}{5.59 \times 10^{-11}}}.$$

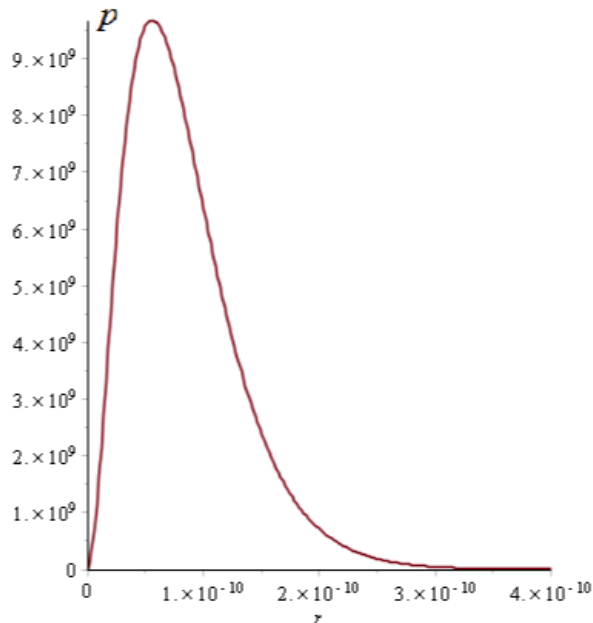
Use Maple to graph the function $p(r)$.

The graph of the density function $p(r)$ is shown below:

Keystrokes:

`plot((4*r/(5.59*10^(-11))^3)*exp(-2*r/(5.59*10^(-11))), r = 0 .. 4*10^(-10));`

Maple result:



(d)

The probability that the electron will be within the sphere of radius r meters centered at the nucleus is

$$P(r) = \int_0^r \frac{4}{a_0^3} s^2 e^{-\frac{2s}{a_0}} ds$$

$$P(r) = \int_0^r \frac{4}{(5.59 \times 10^{-11})^3} s^2 e^{-\frac{2s}{5.59 \times 10^{-11}}} ds$$

When the radius of the sphere $r = 4a_0 \approx 4(5.59 \times 10^{-11})$, the probability is

$$\begin{aligned} P(r) &= \int_0^{4(5.59 \times 10^{-11})} \frac{4}{(5.59 \times 10^{-11})^3} s^2 e^{-\frac{2s}{5.59 \times 10^{-11}}} ds \\ &= \frac{4}{(5.59 \times 10^{-11})^3} \int_0^{4(5.59 \times 10^{-11})} s^2 e^{-\frac{2s}{5.59 \times 10^{-11}}} ds \end{aligned}$$

Use Maple to solve this integral

Keystrokes:

`4*(int(s^2*exp(-2*s/(5.59*10^(-11))), s = 0 .. (4*5.59)*10^(-11))/(5.59*10^(-11))^3;`

Maple result:

$$> \frac{4}{(5.59 \cdot 10^{-11})^3} \cdot \int_0^{4 \cdot 5.59 \cdot 10^{-11}} s^2 \cdot e^{-\frac{2 \cdot s}{5.59 \cdot 10^{-11}}} ds$$

0.9862460327

Therefore the probability that the electron will be within the sphere of radius $r = 4a_0$ meters centered at the nucleus is $\approx \boxed{0.986}$.

(e)

Mean distance of the electron from the nucleus in the ground state of the hydrogen atom is

$$\begin{aligned}\int_{-\infty}^{\infty} r \cdot p(r) dr &= \int_0^{\infty} r p(r) dr = \int_0^{\infty} r \left[\frac{4}{a_0^3} r^2 e^{-\frac{2r}{a_0}} \right] dr && [\text{For } r \geq 0] \\ &= \frac{4}{a_0^3} \int_0^{\infty} \left(r^3 e^{-\frac{2r}{a_0}} \right) dr\end{aligned}$$

Continuation to the above

$$\begin{aligned}&= \frac{4}{a_0^3} \left[-\frac{1}{8} a_0 (3a_0^3 + 6a_0^2 r + 6a_0 r^2 + 4r^3) e^{-\frac{2r}{a_0}} \right]_0^{\infty} \\ &= \frac{4}{a_0^3} \left[\begin{aligned} &-\frac{1}{8} a_0 (3a_0^3 + 6a_0^2 r + 6a_0 r^2 + 4r^3) e^{-\infty} \\ &+ \frac{1}{8} a_0 (3a_0^3 + 6a_0^2 (0) + 6a_0 (0)^2 + 4(0)^3) e^{-0} \end{aligned} \right] \\ &= \frac{4}{a_0^3} \left[\begin{aligned} &-\frac{1}{8} a_0 (3a_0^3 + 6a_0^2 r + 6a_0 r^2 + 4r^3) (0) \\ &+ \frac{1}{8} a_0 (3a_0^3 + 6a_0^2 (0) + 6a_0 (0)^2 + 4(0)^3) (1) \end{aligned} \right] \\ &= \frac{4}{a_0^3} \left[0 + \frac{1}{8} a_0 (3a_0^3) (1) \right] \\ &= \frac{3}{2} a_0\end{aligned}$$

Therefore, the mean distance of the electron from the nucleus in the ground state of the

hydrogen atom is $\boxed{\frac{3}{2} a_0}$.